

Week 8. Logistic Regression 2

This handout is heavily indebted on the following two references:

- Powers, Daniel A. and Yu Xie. 2000. *Statistical Methods for Categorical Data Analysis*. Academic Press.
- Mize, Trenton D. 2019. "Best Practices for Estimating, Interpreting, and Presenting Nonlinear Interaction Effects." *Sociological Science* 6: 81-117.

Install `spost13_ado.pkg` in your machine. Type `net install spost13_ado.pkg` in Stata. `spost13_ado.pkg` is a user-written program developed by Scott Long at Indiana University.

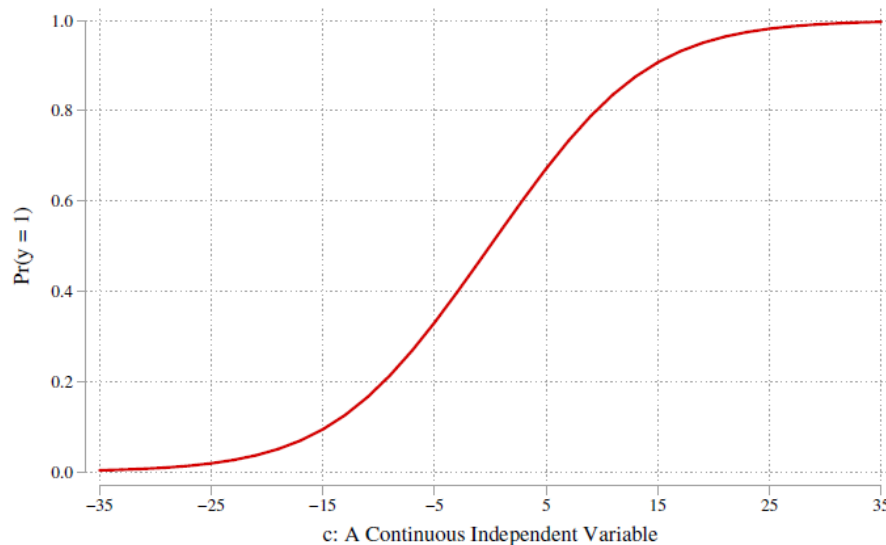
Logit

1. Recall that $\text{logit} = \log \text{odds} = \log \frac{P_A}{1-P_A}$
2. $\text{logit} = b_0 + b_1x_1 + b_2x_2$,
3. $\exp(\log \text{odds}) = \text{odds} = e^{b_0+b_1x_1+b_2x_2}$
4. $\text{odds} = e^{b_0}e^{b_1x_1}e^{b_2x_2}$
5. Thus, as x_1 increases by 1 unit, odds of event A increases by $\exp(b_1)$ times, net of x_2 .
6. $\exp(b_1)$ is called odds ratio.
7. Compared to the reference group (or as x_1 increases by 1 unit), the likelihood of event A is $\exp(b_1)$ times more likely. (Note that here likelihood implies odds.)
8. If you would like to discuss the effect of x on P_A rather than the effect of x on the odds of A, you should report marginal effects.

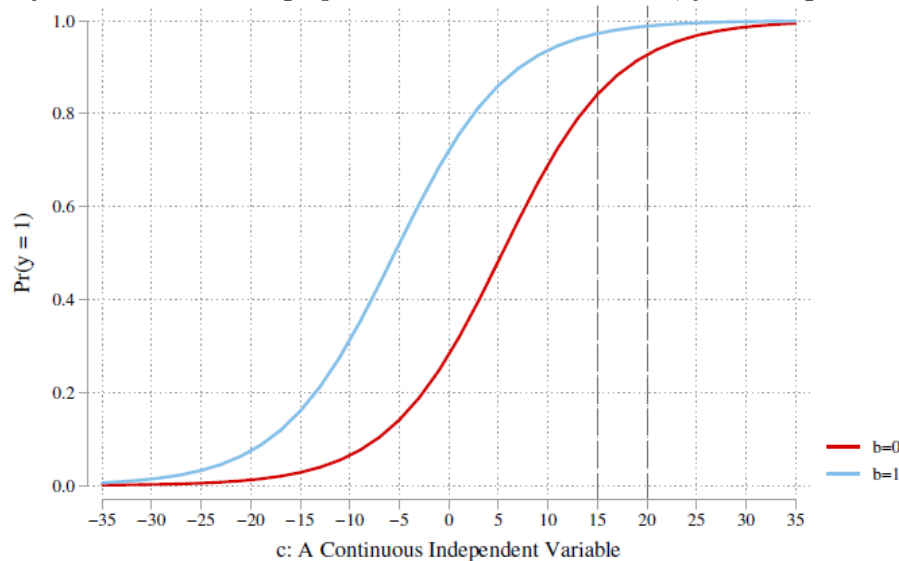
Logit: Interaction effects without interaction terms

1. The effects of independent variables are linear on logit (= log odds).
2. Because all logit functions are multiplicative, the effect of independent variable on P are nonlinear even though there is no interaction effects. That is, the effect of x on P varies across x .
3. Suppose you have the following result:
 $\log \text{odds} = -1 + 0.2c + 2b$
where "c" is a continuous variable and "b" is a binary dummy variable. No interaction terms are added.

4. When you draw the cdf of the continuous variable, you will get the following graph:



5. As you see, the effect of “c” on P varies across “c”. At $c = -25$, the slope of c (or 1st derivative or the effect of c) is small (or very flat), but at $c = 0$, the slope of c is quite steep.
6. If you draw the same graph for both $b = 0$ and $b = 1$, you will get the following graph:



7. The effect of “b” varies across “c”. In the middle of the distribution c, the effect of b is substantial, while the effect of b is tiny at the both ends of the distribution c.
8. The ceiling and flooring effects of logit (and probit) is called “compression.” In essence, the effect of independent variables on P is always interactive even without interaction terms.
9. The effect of “c” differs between $b = 0$ and $b = 1$ depending on “c”. In the middle, the slopes of c are basically the same between $d = 0$ and $d = 1$. At the high end, the slope of c is steep for $b = 0$ while it is flat for $b = 1$.
10. Solution 1: Report marginal effects using the `margins` command which is an average effect.
11. Solution 2: Draw a graph.

12. Solution 3: Compute the marginal effects at different points of interest, using the `margins` and `lincom` (and `mlincom`) command.

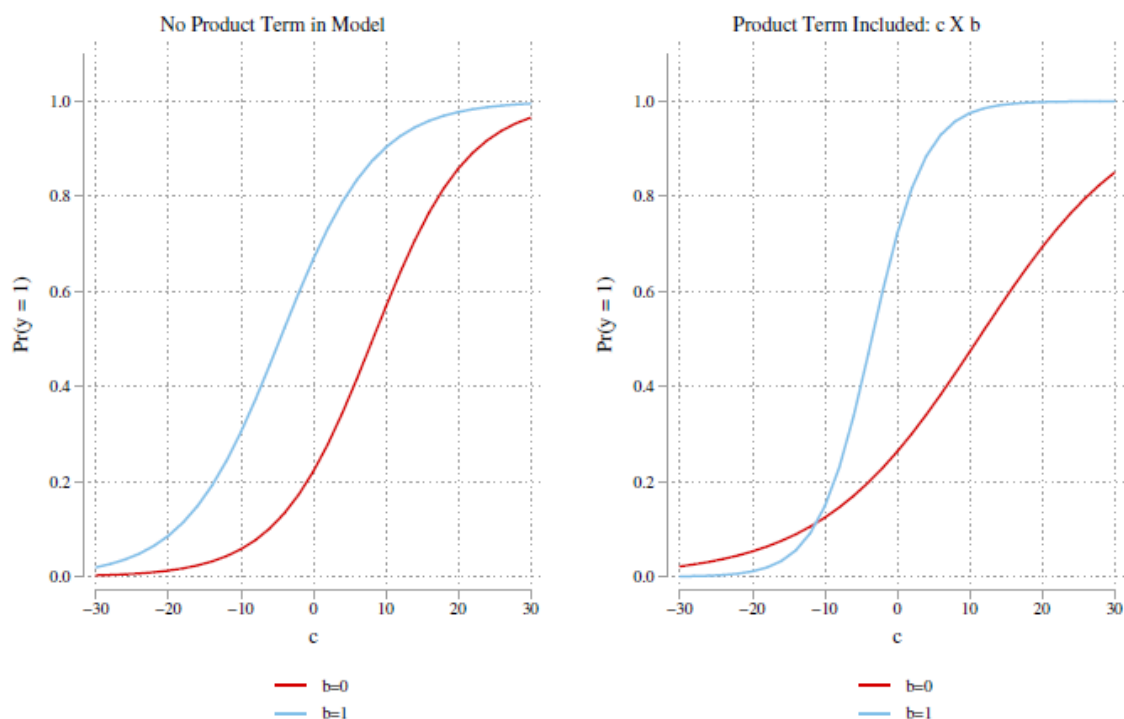
```
margins if b==0, at(c=5)
```

```
margins if b==1, at(c=5)
```

```
mlincom 2-1
```

Logit: Interaction effects with interaction terms

1. Now we add an interaction term on the logit model and get the following coefficients:
 $\log \text{ odds} = -1 + 0.2c + 2b + 0.2(c \times b)$



- 2.
3. The significance of the coefficient estimated for the interaction term does not necessarily indicate whether the interaction is statistically significant. The statistical significance of the interaction term varies across independent variables.
4. Solution 1: The best way to do is to present a graph.
5. Solution 2: Compute the marginal effects at different points of interest, using the `margins` and `lincom` (and `mlincom`) command.

Example:

- Dependent variable: Having a managerial or professional occupation.
- Independent variable: female, education(BA+), racial minority, and age

```
. logit profmanag minority female ba age age2 [pw=perwt]
```

```
Logistic regression                Number of obs   =       22256
                                Wald chi2(5)      =       1980.90
                                Prob > chi2       =        0.0000
Log pseudolikelihood = -429223.45    Pseudo R2      =        0.1288
```

```
-----+-----
      |               Robust
profmanag |               Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
  minority |   -.5967156   .0423213   -14.10   0.000   -.6796638   -.5137674
    female |    .1127462   .0378161     2.98   0.003    .038628   .1868644
         ba |    2.430903   .0561393    43.30   0.000    2.320872   2.540934
        age |   -.1344215   .2073944    -0.65   0.517   -.540907   .272064
       age2 |    .0015181   .0026189     0.58   0.562   -.0036149   .006651
        _cons |    2.093365   4.089812     0.51   0.609   -5.922518   10.10925
-----+-----
```

Non-significant interaction effect: minority * female

```
. logit profmanag i.minority##i.female ba age age2 [pw=perwt]
```

```
Logistic regression                Number of obs   =       22256
                                Wald chi2(6)      =       1985.82
                                Prob > chi2       =        0.0000
Log pseudolikelihood = -429194.32    Pseudo R2      =        0.1288
```

```
-----+-----
      |               Robust
profmanag |               Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
  1.minority |   -.5514066   .0617447    -8.93   0.000   -.6724239   -.4303892
    1.female |    .1387856   .0444581     3.12   0.002    .0516493   .2259219
  minority#female |
    1 1 |   -.0882374   .0844265    -1.05   0.296   -.2537102   .0772354
         ba |    2.431258   .0561342    43.31   0.000    2.321237   2.541279
        age |   -.1343862   .2074167    -0.65   0.517   -.5409154   .2721431
-----+-----
```

```

      age2 |      .0015159      .0026192      0.58      0.563      -.0036176      .0066495
    _cons |      2.082047      4.090049      0.51      0.611      -5.934301      10.0984
-----

```

Significant interaction effect: minority * ba

```
. logit profmanag i.minority##i.ba female age age2 [pw=perwt]
```

```

Logistic regression                                Number of obs   =       22256
                                                    Wald chi2(6)      =       2072.96
                                                    Prob > chi2       =       0.0000
Log pseudolikelihood = -428973.94                Pseudo R2        =       0.1293

```

```

-----
      profmanag |      Coef.      Robust      z      P>|z|      [95% Conf. Interval]
-----+-----
    1.minority |   -.6460647    .046621   -13.86   0.000   -.7374402   -.5546892
         1.ba |    2.300923    .0665198   34.59   0.000    2.170547    2.4313
 minority#ba |
    1 1 |    .3935145    .119472    3.29   0.001    .1593537    .6276754
 female |    .112484    .0378229    2.97   0.003    .0383524    .1866156
    age |   -.1390151    .207292   -0.67   0.502   -.5453     .2672698
   age2 |    .0015759    .0026176    0.60   0.547   -.0035545    .0067062
    _cons |    2.198296    4.08792    0.54   0.591   -5.81388    10.21047
-----

```

Computing marginal effects (or predicted probabilities)

```
. margins minority#ba
```

```

Predictive margins                                Number of obs   =       22256
Model VCE      : Robust

Expression     : Pr(profmanag), predict()

```

```

-----
      |      Delta-method
      |      Margin   Std. Err.      z      P>|z|      [95% Conf. Interval]
-----+-----

```

```

-----+-----
minority#ba |
      0 0 |   .3176858   .0050686   62.68   0.000   .3077515   .3276202
      0 1 |   .822589   .0090821   90.57   0.000   .8047885   .8403895
      1 0 |   .1962522   .0063493   30.91   0.000   .1838079   .2086966
      1 1 |   .7827279   .0154165   50.77   0.000   .7525121   .8129438
-----+-----

```

Testing the difference between majority BA (2nd line) and minority BA (4th line)

```

. mlincom (4-2)

      |   lincom   pvalue      ll      ul
-----+-----
      1 |   -2.188     0.000   -2.339   -2.038

```

Testing the difference between (A) majority BA (2nd line) and minority BA (4th line) and (B) majority non-BA (1st line) and minority non-BA (3rd line).

```

. mlincom (4-2)-(3-1)

      |   lincom   pvalue      ll      ul
-----+-----
      1 |   -3.228     0.000   -3.480   -2.977

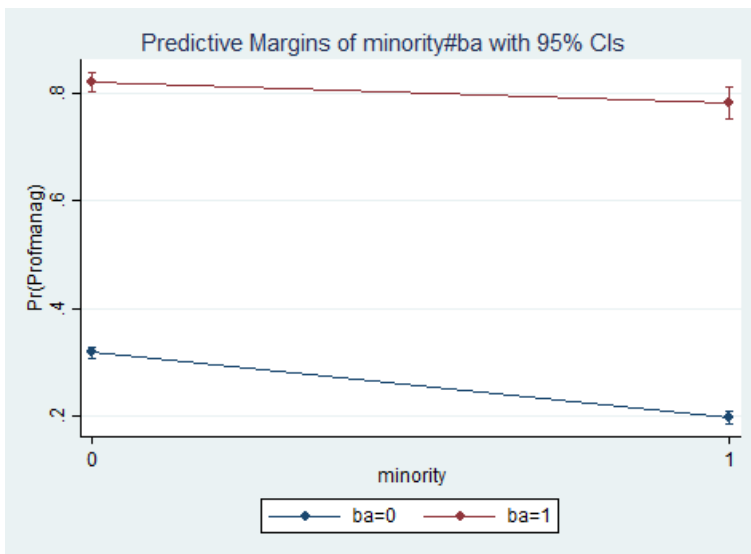
```

Drawing a graph about the interaction effect between minority and BA+

```

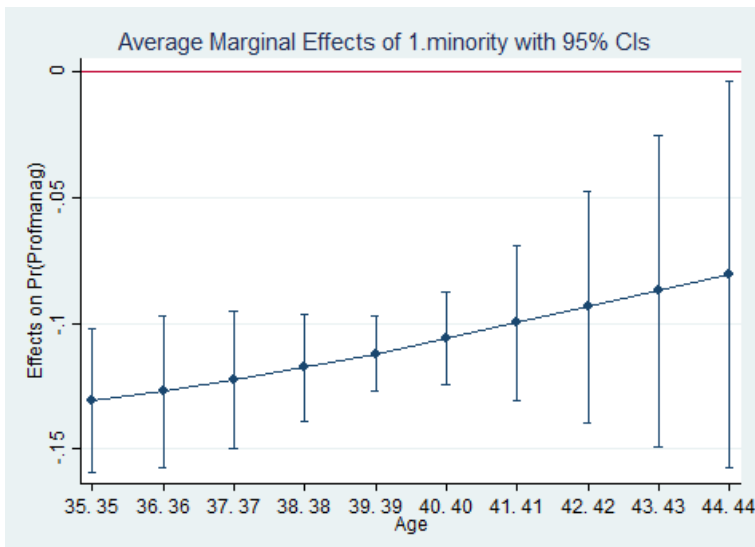
margins minority#ba
marginsplot

```



Drawing a graph on the marginal effect of being minority over age

```
margins, dydx(minority) at(age=(35(1)44))  
marginsplot, yline(0)
```



Multinomial Logit

1. Multinomial Logit models (`mlogit`) can be used when your outcome is multinomial (= 3 or more discrete choices).
2. For logit models, the probability distribution of the event A given the total n follows the binomial distribution. For multinomial logit, the probability distribution of the count Y_j given the total n follows the multinomial distribution.
3. Suppose there are 3 possible outcomes: Y_1 , Y_2 , and Y_3 . (more generally, we can say that there are J possible outcomes Y_1 , Y_2 , ... and Y_J).
4. Then, we can estimate the log odds of event j compared to the reference point J as follows:

$$\log \frac{P_j}{P_J} = \alpha_j + x'\beta_j$$

Note that $\frac{P_j}{P_J}$ is the odds of Y_1 compared to Y_3 . In logit, $\frac{P_A}{1-P_A}$ is the odds of event A compared to event not-A. Thus, logit is a special case of multinomial logit.

If we set Y_3 as a reference event, we can estimate the following two models:

$$\log \frac{P_1}{P_3} = \alpha_1 + x'\beta_1$$

$$\log \frac{P_2}{P_3} = \alpha_2 + x'\beta_2$$

Each coefficient estimated quantifies the change in log odds of Y_j compared to Y_J when x increases by 1 unit. Put differently, $\exp(\beta)$ indicates the odds ratio of Y_j compared to Y_J when x increases by 1 unit. As x increases by 1 unit, the likelihood of Y_j compared to Y_J increases by $\exp(\beta)$ times.

5. The estimation of $\log \frac{P_1}{P_2}$ is not necessary because all coefficients for $\log \frac{P_1}{P_2}$ can be computed from the previous 2 models.

$$\log \frac{P_1}{P_3} - \log \frac{P_2}{P_3} = (\log P_1 - \log P_3) - (\log P_2 - \log P_3) = \log P_1 - \log P_2 = \log \frac{P_1}{P_2}$$

thus,

$$\log \frac{P_1}{P_2} = (\alpha_1 - \alpha_2) + x'(\beta_1 - \beta_2)$$

6. Estimation of the parameters of this model by maximum likelihood proceeds by maximization of the multinomial likelihood with the probabilities P_j viewed as functions of the α_j and β_j parameters.

Multinomial Logit Example

- Dependent Variable: Employment Status = (1) Employed, (2) Unemployed, (3) Not in Labor Force.

```
. mlogit empstat i.edu female minority age age2 [pw=perwt]
```

```

Multinomial logistic regression          Number of obs   =      22256
                                         Wald chi2(16)    =      1421.83
                                         Prob > chi2      =       0.0000
Log pseudolikelihood = -403878.38        Pseudo R2       =       0.0978

```

empstat	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
<hr/>						
1__Employed	(base outcome)					
<hr/>						
2__Unemployed						
edu						
2	-.2816744	.1293254	-2.18	0.029	-.5351475	-.0282013
3	-.5087672	.1303798	-3.90	0.000	-.7643069	-.2532274
4	-1.216158	.143058	-8.50	0.000	-1.496547	-.9357697
5	-1.692009	.1765734	-9.58	0.000	-2.038086	-1.345931
female	.2232413	.0794099	2.81	0.005	.0676006	.3788819
minority	.4576939	.0842396	5.43	0.000	.2925873	.6228005
age	-.6087	.4206603	-1.45	0.148	-1.433179	.2157791
age2	.0076419	.0053096	1.44	0.150	-.0027648	.0180486
_cons	9.629997	8.300133	1.16	0.246	-6.637965	25.89796
<hr/>						
3__Not_in_labor_~e						
edu						
2	-.6377104	.0944152	-6.75	0.000	-.8227608	-.45266
3	-1.040515	.0909344	-11.44	0.000	-1.218743	-.8622865
4	-1.089392	.0942667	-11.56	0.000	-1.274151	-.9046326
5	-1.718559	.1094104	-15.71	0.000	-1.932999	-1.504119
female	2.011937	.0643168	31.28	0.000	1.885878	2.137995
minority	-.0446437	.0568998	-0.78	0.433	-.1561653	.0668779
age	.1359571	.2787946	0.49	0.626	-.4104703	.6823844
age2	-.0017603	.0035227	-0.50	0.617	-.0086647	.0051441
_cons	-4.806532	5.497639	-0.87	0.382	-15.58171	5.968642

Computing the probability of employment by education.

```
. margins edu, atmeans
```

```
Adjusted predictions          Number of obs    =    22,256
Model VCE      : Robust
```

```
1._predict   : Pr(empstat==1__Employed), predict(pr outcome(1))
2._predict   : Pr(empstat==2__Unemployed), predict(pr outcome(2))
3._predict   : Pr(empstat==3__Not_in_labor_force), predict(pr outcome(3))
at           : 1.edu          =    .0918028 (mean)
               2.edu          =    .2230475 (mean)
               3.edu          =    .2969211 (mean)
               4.edu          =    .2437092 (mean)
               5.edu          =    .1445194 (mean)
               0.female       =    .4966023 (mean)
               1.female       =    .5033977 (mean)
               0.minority     =    .6600245 (mean)
               1.minority     =    .3399755 (mean)
               age            =    39.54809 (mean)
               age2           =    1572.167 (mean)
```

		Delta-method				
		Margin	Std. Err.	z	P> z	[95% Conf. Interval]
_predict#edu						
1	1	.7059303	.0142476	49.55	0.000	.6780054 .7338551
1	2	.8024339	.0075069	106.89	0.000	.7877206 .8171472
1	3	.850526	.0055411	153.49	0.000	.8396656 .8613864
1	4	.8806488	.0050144	175.62	0.000	.8708207 .8904769
1	5	.9300099	.0045712	203.45	0.000	.9210505 .9389694
2	1	.0813647	.0080513	10.11	0.000	.0655845 .0971449
2	2	.0697837	.0047661	14.64	0.000	.0604423 .079125
2	3	.0589396	.0038759	15.21	0.000	.0513429 .0665362
2	4	.030082	.0026606	11.31	0.000	.0248672 .0352967
2	5	.0197393	.0027065	7.29	0.000	.0144347 .0250439
3	1	.212705	.0129952	16.37	0.000	.187235 .2381751
3	2	.1277824	.0062557	20.43	0.000	.1155215 .1400433
3	3	.0905344	.0042333	21.39	0.000	.0822373 .0988316
3	4	.0892692	.0043666	20.44	0.000	.0807109 .0978275
3	5	.0502508	.0037632	13.35	0.000	.042875 .0576266

Gender comparison across the probabilities of three outcomes.

```
. margins edu, predict() over(i.female) atmeans
```

Adjusted predictions	Number of obs	=	22,256
Model VCE : Robust			

```
over      : female
1._predict : Pr(empstat==1__Employed), predict(pr outcome(1))
2._predict : Pr(empstat==2__Unemployed), predict(pr outcome(2))
3._predict : Pr(empstat==3__Not_in_labor_force), predict(pr outcome(3))
at        : 0.female
```

1.edu	=	.1029445 (mean)
2.edu	=	.2354108 (mean)
3.edu	=	.2854081 (mean)
4.edu	=	.235196 (mean)
5.edu	=	.1410406 (mean)
female	=	0
0.minority	=	.6620612 (mean)
1.minority	=	.3379388 (mean)
age	=	39.54567 (mean)
age2	=	1572.055 (mean)
1.female		
1.edu	=	.0808115 (mean)
2.edu	=	.2108511 (mean)
3.edu	=	.3082787 (mean)
4.edu	=	.2521074 (mean)
5.edu	=	.1479512 (mean)
female	=	1
0.minority	=	.6580153 (mean)
1.minority	=	.3419847 (mean)
age	=	39.55048 (mean)
age2	=	1572.278 (mean)

```

2 1 4 | .0290641 .002782 10.45 0.000 .0236115 .0345168
2 1 5 | .0202658 .0029099 6.96 0.000 .0145626 .0259691
3 0 1 | .0902598 .0073401 12.30 0.000 .0758734 .1046462
3 0 2 | .0509319 .0035917 14.18 0.000 .0438923 .0579714
3 0 3 | .0351265 .00243 14.46 0.000 .0303637 .0398893
3 0 4 | .0344926 .0023514 14.67 0.000 .029884 .0391012
3 0 5 | .0188947 .0017217 10.97 0.000 .0155203 .0222691
3 1 1 | .4202851 .0190743 22.03 0.000 .3829002 .4576699
3 1 2 | .2827419 .0108342 26.10 0.000 .2615073 .3039765
3 1 3 | .2115603 .0078637 26.90 0.000 .1961478 .2269728
3 1 4 | .2096072 .0088578 23.66 0.000 .1922463 .2269682
3 1 5 | .12538 .0084167 14.90 0.000 .1088836 .1418765

```

```
-----
. margins i.edu, dydx(i.female) atmeans
```

```
Conditional marginal effects          Number of obs    =    22,256
Model VCE      : Robust
```

```
dy/dx w.r.t. : 1.female
```

```
1._predict : Pr(empstat==1__Employed), predict(pr outcome(1))
2._predict : Pr(empstat==2__Unemployed), predict(pr outcome(2))
3._predict : Pr(empstat==3__Not_in_labor_force), predict(pr outcome(3))
```

```
at      : 1.edu      = .0918028 (mean)
          2.edu      = .2230475 (mean)
          3.edu      = .2969211 (mean)
          4.edu      = .2437092 (mean)
          5.edu      = .1445194 (mean)
          0.female   = .4966023 (mean)
          1.female   = .5033977 (mean)
          0.minority = .6600245 (mean)
          1.minority = .3399755 (mean)
          age        = 39.54809 (mean)
          age2       = 1572.167 (mean)
```

```
-----
          |              Delta-method
          |      dy/dx   Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----
0.female | (base outcome)
-----+-----
1.female |
_predict#edu |
1 1 | -.3111994 .0140035 -22.22 0.000  -.3386457  -.2837532
1 2 | -.2268817 .0098033 -23.14 0.000  -.2460957  -.2076676
1 3 | -.1768153 .0077511 -22.81 0.000  -.1920073  -.1616233
1 4 | -.1755627 .0080303 -21.86 0.000  -.1913018  -.1598236
1 5 | -.1084769 .0072672 -14.93 0.000  -.1227204  -.0942334
2 1 | -.0188237 .0058193  -3.23 0.001  -.0302294  -.0074181
2 2 | -.0049249 .0049082  -1.00 0.316  -.0145447  .0046949
2 3 | .0003853 .0042242   0.09 0.927  -.0078941  .0086646
2 4 | .0004538 .0022263   0.20 0.838  -.0039097  .0048173
2 5 | .0019959 .0015355   1.30 0.194  -.0010137  .0050054

```

3 1		.3300231	.0147675	22.35	0.000	.3010794	.3589669
3 2		.2318066	.0094055	24.65	0.000	.2133722	.2502409
3 3		.17643	.0070254	25.11	0.000	.1626606	.1901995
3 4		.1751088	.0079644	21.99	0.000	.1594989	.1907188
3 5		.1064811	.0072558	14.68	0.000	.09226	.1207022

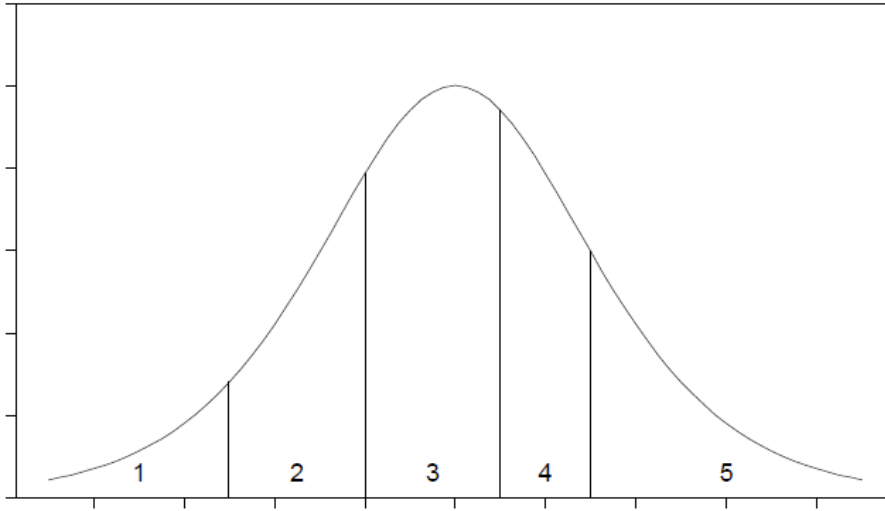
 Note: dy/dx for factor levels is the discrete change from the base level.

If you would like to change the base outcome to (2) Unemployed, estimate the following command.

```
mlogit empstat i.edu female minority age age2 mar[pw=perwt], baseoutcome(2)
```

Ordinal Logit or Ordered Logit

1. Ordered Logit models (`ologit`) can be used when your outcome variable is ordered (for example, high=3, medium=2, and low=1; for another example, strongly agree=4, somewhat agree=3, somewhat disagree=2, strongly disagree=1).
2. These models can also be interpreted in terms of a latent variable. Specifically, suppose that the manifest response Y_i results from grouping an underlying continuous variable Y_i^* using cut-points $\theta_1 < \theta_2 \dots < \theta_{J-1}$, so that Y_i takes the value 1 if Y_i^* is below θ_1 , the value 2 if Y_i^* is between θ_1 and θ_2 , and so on, taking the value J if Y_i^* is above θ_{J-1} .



3. Thus, Ordered Logit estimates the effect of independent variables as a function of cumulative probabilities. Because it is a cumulative probability function, Ordered Logit estimates one coefficient for each independent variable and compute
4. The cumulative probability C_{ij} for individual i up to response level j ,

$$C_{ij} = Pr(y_j \leq j) = Pr(y_i \leq j|x) = Pr(x'\beta + \epsilon \leq \theta_j),$$

Rearrange terms, we find that:

$$C_{ij} = Pr(\epsilon \leq \theta_j - x'\beta) = F(\theta_j - x'\beta)$$

Thus, the probability of $y_j = j$ given x is,

$$Pr(y_i = j|X_i) = \begin{cases} F(\theta_1 - x'\beta), & j = 1 \\ F(\theta_j - x'\beta) - F(\theta_{j-1} - x'\beta), & 1 < j \leq J - 1 \\ 1 - F(\theta_{J-1} - x'\beta), & j = J \end{cases}$$

In Stata, θ_j are cut points. Note that only one set of β 's are estimated.

5. The probability of C_{ij} can be estimated as follows:

$$C_{ij} = \frac{\exp(\theta_j + x'\beta)}{1 + \exp(\theta_j + x'\beta)}$$

6. Interpretation: Conditional on the other covariates, the odds that Y_j is less than or equal to a given level j versus greater than j is estimated to be $\exp(\beta)$ times greater as x increases by 1 unit.

Ordered logit: dep variable = Subjective Well-being (5= high, 1=low)

```
. ologit SWB SSS c.age##c.age female yrsch [pw=weight]
```

```
Iteration 0:  log pseudolikelihood = -13331.542
Iteration 1:  log pseudolikelihood = -12972.818
Iteration 2:  log pseudolikelihood = -12969.055
Iteration 3:  log pseudolikelihood = -12969.053
Iteration 4:  log pseudolikelihood = -12969.053
```

```
Ordered logistic regression          Number of obs    =    11,001
                                   Wald chi2(5)         =    535.31
                                   Prob > chi2           =    0.0000
Log pseudolikelihood = -12969.053   Pseudo R2         =    0.0272
```

	SWB	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
SSS		.288463	.0133989	21.53	0.000	.2622015	.3147244
age		-.0197745	.0068535	-2.89	0.004	-.0332072	-.0063418
c.age#c.age		.0001688	.0000676	2.50	0.013	.0000362	.0003013
female		-.0299368	.0400124	-0.75	0.454	-.1083596	.048486
yrsch		.0028191	.0050849	0.55	0.579	-.0071471	.0127854
/cut1		-3.373698	.2013347			-3.768307	-2.979089
/cut2		-1.68202	.1899033			-2.054223	-1.309816
/cut3		-.2022073	.1863877			-.5675205	.1631058
/cut4		2.366788	.1875696			1.999158	2.734417

Estimating the probabilities of 5 Likert-scale outcomes.

```
. margins
```

```
Predictive margins          Number of obs    =    11,001
Model VCE      : Robust
```

```

1._predict : Pr(SWB==1), predict(pr outcome(1))
2._predict : Pr(SWB==2), predict(pr outcome(2))
3._predict : Pr(SWB==3), predict(pr outcome(3))
4._predict : Pr(SWB==4), predict(pr outcome(4))
5._predict : Pr(SWB==5), predict(pr outcome(5))

```

		Delta-method					
		Margin	Std. Err.	z	P> z	[95% Conf. Interval]	

_predict							
1		.0168795	.0013263	12.73	0.000	.01428	.0194789
2		.0668575	.0025816	25.90	0.000	.0617977	.0719173
3		.1920383	.0039601	48.49	0.000	.1842766	.1998001
4		.5374719	.005133	104.71	0.000	.5274114	.5475323
5		.1867528	.0039678	47.07	0.000	.1789761	.1945296

Comparing the probabilities of 5 Likert-scale outcomes by Subjective Social Standing

```

. margins, at(SSS=(1(3)10))

```

```

Predictive margins                                Number of obs    =    11,001
Model VCE      : Robust

```

```

1._predict : Pr(SWB==1), predict(pr outcome(1))
2._predict : Pr(SWB==2), predict(pr outcome(2))
3._predict : Pr(SWB==3), predict(pr outcome(3))
4._predict : Pr(SWB==4), predict(pr outcome(4))
5._predict : Pr(SWB==5), predict(pr outcome(5))

```

```

1._at      : SSS          =          1
2._at      : SSS          =          4
3._at      : SSS          =          7
4._at      : SSS          =         10

```

		Delta-method					
		Margin	Std. Err.	z	P> z	[95% Conf. Interval]	

_predict#_at							
1 1		.0405484	.0035615	11.39	0.000	.0335681	.0475288
1 2		.0174791	.0013727	12.73	0.000	.0147887	.0201694
1 3		.0074324	.0006443	11.54	0.000	.0061696	.0086951

1 4		.0031418	.0003445	9.12	0.000	.0024667	.003817
2 1		.1459302	.0071279	20.47	0.000	.1319598	.1599006
2 2		.0705568	.0027353	25.79	0.000	.0651958	.0759179
2 3		.0316223	.0016948	18.66	0.000	.0283005	.034944
2 4		.0136784	.001178	11.61	0.000	.0113695	.0159872
3 1		.3147249	.0074074	42.49	0.000	.3002068	.3292431
3 2		.2094484	.0043884	47.73	0.000	.2008473	.2180494
3 3		.1123264	.0041648	26.97	0.000	.1041636	.1204891
3 4		.0530482	.0039051	13.58	0.000	.0453944	.060702
4 1		.4277687	.0095175	44.95	0.000	.4091147	.4464227
4 2		.5488553	.0052543	104.46	0.000	.538557	.5591537
4 3		.5474551	.0058007	94.38	0.000	.5360859	.5588242
4 4		.4245533	.0136385	31.13	0.000	.3978223	.4512842
5 1		.0710278	.0039199	18.12	0.000	.0633448	.0787107
5 2		.1536604	.0038903	39.50	0.000	.1460356	.1612852
5 3		.3011639	.0081255	37.06	0.000	.2852381	.3170897
5 4		.5055784	.0182223	27.75	0.000	.4698633	.5412934
